# Characteristic Vector Analysis of $S_{\mathrm{N}} 1$ Kinetic Data Influenced by Four Variables $\dagger$ 

Romuald I. Zalewski * and Zdzisława Geltz<br>Department of General Chemistry, Academy of Economics, Poznań, Poland

Characteristic vector analysis has been used to explain the variability in the kinetic data of 1260 solvolysis reactions caused by four variables. Two characteristic vectors $V_{1}$ and $V_{2}$ describe the variability with sufficient accuracy. $V_{1}$ and $V_{2}$ are linearly related to $\sigma_{0}$ substituent constants. $S_{1}{ }^{\text {a }}$ and $S_{2}{ }^{\prime}$ (scalar values after rotation of axes) are linearly dependent on the solvent polarity parameter $Y$ and the reciprocal of the temperature, respectively.

An estimation of the degree of efficiency of various variables that simultaneously affect the reactivity of organic molecules has been based on the principle of additivity and independence. ${ }^{1}$

The additivity approach could be misleading because of the existence of perturbation among the various types of interaction. The logical consequence was to include cross-terms in the correlation analysis. Such mathematical models were designed to determine quantitatively the contribution of interacting effects in complicated physico-chemical data sets. ${ }^{2}$

There are few examples of the application of such models to the study of reactions influenced by a large number of independent factors. However, Pivovarov et al. ${ }^{3}$ published a large set of kinetic data on solvolysis of $\beta$-dinitroalkylarylamines. The reaction takes place via an $S_{\mathrm{N}} 1$ mechanism, and variation of the reaction rates is influenced by a change in two substituents, temperature and polarity of the medium, and could be expressed by a complicated equation consisting of 16 variable terms and cross-terms of different complexity.

The practical application of such a procedure to explain experimental data or for the prediction of reaction rates for unknown compounds is very difficult, if not impossible.

This paper is an attempt at a more direct approach to study data sets that are influenced by a large number of independent variables. This approach is based on the characteristic vector analysis of a large data set for a solvolysis reaction. The advantage of this approach is that it is not based on a specific theory relating the experimental data to the variables, but on a general theory of the behaviour of data in similar processes.

Design of the Study.-In the present paper we studied the processes involved in the solvolysis reactions of $\beta$-dinitroalkylamines, of general formula $\mathrm{R}^{1} \mathrm{C}_{6} \mathrm{H}_{4} \mathrm{NHCH}_{2} \mathrm{C}\left(\mathrm{NO}_{2}\right)_{2} \mathrm{C}_{6} \mathrm{H}_{4} \mathrm{R}^{2}$. The reaction rate is affected by four types of variables: $\mathrm{R}^{1}, \mathrm{R}^{2}$, temperature, and solvent composition. The reaction mechanism is $S_{\mathrm{N}} 1$ and is the same for all compounds. This is contrary to the recent results of Albano and Wold ${ }^{4}$ who studied various $S_{\mathrm{N}} 1$ and $S_{\mathrm{N}} 2$ reactions.

In the present study 42 compounds (objects, $i$ ) were used. Six series of compounds for reaction in which $\mathrm{R}^{1}=p$ - Me (1), H (2), $p-\mathrm{Br}$ (3), $m-\mathrm{Br}$ (4), $p-\mathrm{CO}_{2} \mathrm{Me}$ (5), and $m-\mathrm{NO}_{2}$ (6) were synthesized, each having seven members with various substituents, $\mathrm{R}^{2}=p-\mathrm{Me}, p-\mathrm{OMe}, \mathrm{H}, p-\mathrm{Br}, m-\mathrm{Cl}, m-\mathrm{NO}_{2}$, and $p-\mathrm{NO}_{2}$. The reaction rate for each object was measured under 30 different sets of conditions (variables, $m$ ). Reactions at five different temperatures, $10,15,25,35$, and $55^{\circ} \mathrm{C}$, were carried out in six different media containing $0,30,50,70,80$, and $100 \% \mathrm{v} / \mathrm{v}$ methanol.

All the kinetic data ( 1260 values) have been arranged in a

[^0]complete experimental data matrix of format $42 \times 30$. In each row, the kinetic data were arranged for 42 combinations of $\mathrm{R}^{1}$ and $\mathbf{R}^{2}$ (objects) under constant conditions (variables), i.e., temperature and medium. In each column the kinetic data were arranged for one compound at 30 combinations of variables. The schematic arrangement is presented in the Scheme.

According to the specificity of characteristic vector analysis, by selecting processes of known types it is possible to use a mathematical model for: (i) classification of objects as having or not having data consistent with the model; (ii) classification of variables; and (iii) classification of processes.

We will deal with points (i) and (ii) only.
Mathematical Description of Characteristic Vector Analysis. - Characteristic vector analysis (c.v.a.) has been used in this laboratory for various purposes and has been described previously. ${ }^{5,6}$ A brief description only will be given here. Reaction rates, $\log k$ for $i=1,2 \ldots n$ objects and $m=$ $1,2 \ldots k$ variables, form a data matrix $\log \mathrm{k}_{l, m}$. The experimental data matrix could be approximated by the c.v.a. equa-
$\log k_{m, t}=\log k_{m, l}+V_{1, t} S_{m, 1}+\ldots+V_{k, t} S_{m, k}+E_{l, m}(1)$
tion (1), where characteristic vectors $V$ are specific to the variations of objects $i$ and $S$ describes their magnitudes which must be added to the mean data vector, $\log k$, in order to reconstitute the experimental data set.

The number of characteristic vectors will be equal to or less than $n$ and in general is much less than $n$. The importance of the particular vector is measured by the percentage of the total variability explained by this vector, $\% \mathrm{TV}$, and by the residual variance of the difference between experimental and reconstituted data, $W_{E^{2}}$. The calculation is performed until $\% \mathrm{TV}=$ 100 or just using the declared number of characteristic vectors.

## Results and Discussion

The experimental data matrix $\log k_{m, l}$ can be approximated using the first characteristic vector $V_{1, l}$ with the total variability of $97.18 \%$. The application of the second characteristic vector will improve the $\%$ TV to $99.25 \%$.

The variance of unexplained residual $E_{t, m}$ after application of the first and second characteristic vectors calculated for all objects and variables is $0.001-0.035$ and $<0.01$, respectively. Thus the standard deviation of reconstituted $\log k$ is acceptable when both characteristic vectors are used to explain the variability among all kinetic data, equation (2). $V_{1, \iota}, V_{2, t}$, and

$$
\begin{equation*}
\log k_{m, l}=\log k_{m, l}+V_{1, l} S_{m, 1}+V_{2, l} S_{m, 2}+E_{m, l} \tag{2}
\end{equation*}
$$

$\log \bar{k}_{m, l}$ are the 42 set of numbers reported in Table 1. In the


Scheme.

Table 1. Values of $V_{1}, V_{2}$, and $\log k$

| Series | $\underset{R^{1}}{\text { Substituent }}$ | Calculated values | Substituent $\mathbf{R}^{\mathbf{2}}$ |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | p-Me | $p$-OMe | H | $p-\mathrm{Br}$ | $m-\mathrm{Cl}$ | $m-\mathrm{NO}_{2}$ | $p-\mathrm{NO}_{2}$ |
| 1 | $p-\mathrm{Me}$ | $V_{1}$ | 6.4714 | 6.3340 | 5.7885 | 4.6014 | 4.1016 | 2.6006 | 2.0479 |
|  |  | $V_{2}$ | 1.7690 | 1.7018 | 1.4369 | 0.8567 | 0.6127 | -0.1157 | -0.3966 |
|  |  | $\log k$ | -1.8170 | -1.7740 | -1.6021 | -1.2302 | -1.0724 | -0.6007 | -0.4297 |
| 2 | H | $V_{1}$ | 6.4733 | 6.3516 | 5.8662 | 4.8163 | 4.3707 | 3.0375 | 2.5532 |
|  |  | $V_{2}$ | 1.1355 | 1.0927 | 0.9206 | 0.5522 | 0.3942 | -0.0757 | -0.2472 |
|  |  | $\log k$ | -2.2992 | -2.2549 | -2.0771 | -1.6912 | -1.5282 | -1.0389 | -0.8605 |
| 3 | $p-\mathrm{Br}$ | $V_{1}$ | 6.4741 | 6.3750 | 5.9734 | 5.1079 | 4.7423 | 3.6424 | 3.2437 |
|  |  | $V_{2}$ | 0.2696 | 0.2590 | 0.2183 | 0.1296 | 0.0982 | -0.0171 | -0.0569 |
|  |  | $\log k$ | -2.9571 | -2.9101 | -2.7244 | -2.3208 | -2.1499 | -1.6381 | -1.4517 |
| 4 | $m-\mathrm{Br}$ | $V_{1}$ | 6.4754 | 6.3896 | 6.0473 | 5.3098 | 4.9964 | 4.0582 | 3.7168 |
|  |  | $V_{2}$ | -0.3239 | -0.3114 | -0.2625 | -0.1564 | -0.1114 | 0.0229 | 0.0724 |
|  |  | $\log k$ | -3.4087 | -3.3601 | -3.1677 | -2.7522 | -2.5763 | -2.0487 | -1.8572 |
| 5 | $p-\mathrm{CO}_{2} \mathrm{Me}$ | $V_{1}$ | 6.4783 | 6.4153 | 6.1644 | 5.6217 | 5.3873 | 4.6902 | 4.4439 |
|  |  | $V_{2}$ | -1.2361 | -1.1902 | $-1.0039$ | -0.5921 | -0.4293 | 0.1041 | 0.2691 |
|  |  | $\log k$ | -4.1021 | -4.0523 | -3.8517 | -3.4196 | -3.2334 | -2.6837 | -2.4814 |
| 6 | $m-\mathrm{NO}_{2}$ |  | 6.4787 | 6.4175 | 6.1981 | 5.7322 | 5.5046 | 4.8898 | 4.6713 |
|  |  | $V_{2}$ | -1.5119 | -1.4581 | -1.2269 | -0.8098 | -0.5267 | 0.1030 | 0.3434 |
|  |  | $\log k$ | -4.3122 | -4.2609 | -4.0583 | -3.6273 | -3.4321 | -2.8731 | -2.6685 |

first line values for seven $\mathbf{R}^{2}$ substituents are reported, substituent $\mathbf{R}^{1}$ being held constant (series 1 ). The other lines contain data for series $2-6$. The linear regression between $V_{1, t}$ for series $1-6$ yields a correlation coefficient of $r>0.999$, the data for series 2 being treated as the reference $\left(\mathrm{R}^{1}=H\right)$. The same holds for $V_{2, l}$. In addition $V_{1, l}$ linearly relates to $\sigma_{0}$ substituent constants (Figure 1) with a very high correlation coefficient. $V_{2, i}$ behaves in a similar way; however, the parameters of the linear equations are different. There is no doubt that both vectors reflect the action of substituents $\mathbf{R}^{1}$ and $\mathbf{R}^{2}$ on reactivity.

The explanation of $S_{m, 1}$ and $S_{m, 2}$ values is more complicated. The plot of $S_{2}$ versus $S_{1}$ presented in Figure 2(a) indicates a complex situation. The points form straight lines. Six lines represent the data for different solvent composition (negative slope) and the other five lines represent the data for different temperatures (positive slope). However, the distribution of the data is asymmetrical. For example, $S_{2}$ values for $100 \%$ methanol overlap with the $S_{2}$ values for 80 and $70 \%$ methanol. The $S_{1}$ values behave in the same way. To avoid difficulties we rotated the $S_{1}$ and $S_{2}$ axis by $54^{\circ}$. The new plot $S_{m, 1^{\prime}}$ versus


Figure 1. Values of $V_{1, i}$ for series $2\left(\mathrm{R}^{1}=\mathrm{H}\right)$ versus $\sigma_{0}$ substituent constants of $\mathrm{R}^{2}$


Figure 2. Plot of $S_{2}$ versus $S_{1}$ : (a) indicates original data and (b) indicates data after rotation

Table 2. Values of $S_{1}$ and $S_{2}$ for all solvent-temperature combinations

| Solvent | Calculated values | Temperature (K) |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 283 | 288 | 298 | 308 | 328 |
| $\mathrm{H}_{2} \mathrm{O}$ | $S_{1}$ | -0.0106 | 0.0380 | 0.1311 | 0.2178 | 0.3758 |
|  | $S_{2}$ | 0.3701 | 0.3199 | 0.2259 | 0.1368 | -0.0245 |
| MeOH, 30\% | $S_{1}$ | -0.0802 | -0.0286 | 0.0672 | 0.1571 | 0.3202 |
|  | $S_{2}$ | 0.2756 | 0.2223 | 0.1379 | 0.0571 | -0.0879 |
| MeOH, 50\% | $S_{1}$ | -0.1342 | -0.0823 | 0.0160 | 0.1080 | 0.2752 |
|  | $S_{2}$ | 0.1845 | 0.1433 | 0.0658 | <0.0066 | -0.1376 |
| MeOH, 70\% | $S_{1}$ | -0.2038 | -0.1504 | -0.0490 | 0.0459 | 0.2183 |
|  | $S_{2}$ | 0.0786 | 0.0432 | -0.0243 | -0.0875 | -0.2022 |
| MeOH, 80\% | $S_{1}$ | -0.2505 | -0.1959 | -0.0925 | 0.0043 | 0.1802 |
|  | $S_{2}$ | 0.0093 | -0.0250 | -0.0847 | -0.1417 | -0.2445 |
| MeOH, 100\% | $S_{1}$ | -0.3637 | -0.3067 | -0.1981 | -0.0966 | 0.0878 |
|  | $S_{2}$ | -0.1629 | -0.1862 | -0.2314 | -0.2732 | -0.3503 |

Table 3. Parameters of regression for $S_{1}{ }^{\prime}$ and $S_{2}{ }^{\prime}$

| $T /{ }^{\circ} \mathrm{C}$ | $S_{1}{ }^{\prime}=a Y+b$ | $r$ | $\mathrm{MeOH}(\%)$ | $S_{2}{ }^{\prime}=c / T+d$ | $r$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 10 | $0.1879 Y-0.1295$ | 0.9959 | 0 | $-1121.811 / T+3.7404$ | 0.9999 |
| 15 | $0.1853 Y-0.1313$ | 0.9988 | 30 | $-1104.211 / T+3.6795$ | 0.9999 |
| 25 | $0.1705 Y-0.1208$ | 0.9987 | 50 | $-1073.411 / T+3.5769$ | 0.9999 |
| 35 | $0.1339 Y-0.1022$ | 0.9899 | 70 | $-1045.431 / T+3.4827$ | 0.9999 |
| 55 | $0.1317 Y-0.0926$ | 0.9987 | 80 | $-1026.481 / T+3.4186$ | 0.9999 |
|  |  |  | 100 | $-982.811 / T+3.2713$ | 0.9999 |

$S_{m .2}{ }^{\prime}$ is presented by Figure 2(b). The distribution of points along $S_{1}{ }^{\prime}$ and $S_{2}{ }^{\prime}$ is then symmetrical and their ordinates are expressed by equations (3) and (4).

$$
\begin{align*}
& S_{1}^{\prime}=S_{1} \cos \alpha+S_{2} \sin \alpha  \tag{3}\\
& S_{2}^{\prime}=S_{2} \cos \alpha-S_{1} \sin \alpha \tag{4}
\end{align*}
$$

$S_{m, 1^{\prime}}$ represents the effect of the solvent composition and, for a particular temperature, is linearly related to the polarity scale, $Y .^{7}$ At $298 K S_{m, 2}{ }^{\prime}$ is constant and is -0.025 . $S_{m, 2}{ }^{\prime}$
represents the effect of temperature and for a particular solvent composition is linearly related to the reciprocal of the temperature (Table 2). The parameters of regression are given in Table 3.

## References

1 V. Palm, ' Fundamentals of Quantitative Theory of Organic Reactions,' Chimia, Moscow, 1967.
2 S. I. Miller, J. Am. Chem. Soc., 1959, 81, 101.

3 S. A. Pivovarov, V. P. Selivanov, and B. V. Gidaspov, Org. React., 1975, 12, 305.
4 C. Albano and S. Wold, J. Chem. Soc., Perkin Trans. 2, 1980, 1447.

5 R. I. Zalewski, J. Chem. Soc., Perkin Trans. 2, 1979, 1637.
6 R. I. Zalewski, A. J. Sarkice, and Z. Geltz, J. Chem. Soc., Perkin Trans. 2, 1983, 1059.

7 A. H. Feinberg and S. Winstein, J. Am. Chem. Soc., 1957, 79, 1597, 1602, 1608; S. Winstein, H. Feinberg, and E. Grunwald, ibid., p. 4146.


[^0]:    $\dagger$ Presented at a Conference on Correlation Analysis in Organic Chemistry, Hull, 1982.

